

Opinion

Multiplicity corrections in life sciences: challenges and consequences

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Introduction

The issue of nonreproducibility has become a pervasive challenge across life sciences, raising concerns about the reliability of research findings. It has been argued that current practices may lead to a substantial proportion of false-positive results in published research findings under certain conditions [1]. This concern regarding inflated error rates and decreased reproducibility, resulting from probability pyramiding, was highlighted decades ago, emphasizing the need for independent replication in research [2]. While psychology and social sciences have faced notable criticism for low reproducibility rates, medical research is not exempt. Highly cited research studies often report substantial initial effects that are contradicted or diminished in follow-up investigations [3]. For instance, the large-scale “Reproducibility Project: Cancer Biology” found consistent results in just 26% of attempted replications, with replication effect sizes averaging 85% smaller than initially reported [4].

Central to the reproducibility crisis is the growing complexity of modern research designs, especially in clinical trials and epidemiology. Trials often feature multiple endpoints, treatment arms, or exploratory analyses, whereas epidemiological studies rely on large datasets with numerous variables, each introducing the potential for multiple comparisons. Without proper statistical adjustments, this multiplicity increases the risk of false positives, thereby undermining the reliability of the findings. These risks are compounded by academia’s “publish or perish” culture, which favors novel, significant results over rigorous, reproducible work. Selective reporting, p-hacking, and inadequate multiple-hypothesis adjustments distort the evidence base, waste resources, and can lead to misguided clinical decisions [5].

Multiplicity-related issues are particularly concerning in pharmaceutical development, where late-stage trial failures result in significant financial and time losses [3]. Addressing these challenges requires robust statistical practices, prespecified analytical plans, and appropriate corrections for multiple comparisons. Despite long-standing calls for methodological

rigor, the application of these adjustments remains inconsistent, even in high-impact journals.

This Opinion paper examines the prevalence and impact of multiplicity-related errors in clinical and epidemiological research. By identifying key gaps in current practices, we propose actionable solutions to enhance the reproducibility and reliability of scientific findings.

Multiplicity in clinical trials

Multiplicity is an inherent challenge in most randomized clinical trials (RCTs). Sponsors often aim to maximize insights from a single study, leading to complex designs that test multiple doses, regimens, endpoints, treatment arms, predictors, subgroups, or populations, often repeated over time. As John Tukey noted in his 1977 paper, “once multiple questions are to be asked, there will be pressures, some ethical in nature, to concentrate upon those questions for which the results appear most favorable” [6].

While multi-arm trials enhance efficiency by reducing the sample size required compared to separate trials, they also amplify the risk of false positives. For instance, in a trial with multiple comparisons, an unadjusted significance level of $\alpha = 0.05$ can lead to substantial inflation of the type I error—a study testing five hypotheses has approximately a 23% probability of producing at least one false positive. A systematic review of 1351 randomized trials published in PubMed in a single month of 2012 revealed that 79% were parallel-group trials, with 14% involving three arms and 7% having four or more arms [7]. Despite the increasing prevalence of multi-arm RCTs, inadequate correction for multiplicity remains a persistent issue, undermining the validity of reported findings.

Prespecifications of analyses to prevent p-hacking

Addressing multiplicity during the design phase of clinical trials is essential to avoid bias, particularly p-hacking, where

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investigators selectively adopt analysis strategies based on preliminary data review [8]. Established guidelines such as the International Council for Harmonisation (ICH) E9 guideline, *Statistical Principles for Clinical Trials*, or the SPIRIT (Standard Protocol Items: Recommendations for Interventional Trials) stress the need to prespecify statistical methods but provide limited guidance on how analyses should be conducted. To enhance transparency and prevent p-hacking, Kahan *et al.* [8] proposed the pre-SPEC framework, which includes (i) prespecifying analyses before recruitment, (ii) defining a single primary analysis strategy, (iii) creating detailed plans for each analysis, (iv) providing enough detail for independent replication, and (v) ensuring adaptive strategies follow predetermined decisions. Researchers can ensure rigorous, transparent, and reproducible trial analyses by incorporating these measures.

Confirmatory versus exploratory studies

The distinction between confirmatory and exploratory studies should be established early. Confirmatory studies focus on a small set of predefined primary outcomes, with a statistical analysis plan that accounts for family-wise error rate (FWER), *P*-value adjustments, and transparent reporting [9]. On the other hand, exploratory studies are hypothesis-generating and may not require stringent multiple-testing adjustments provided that findings are clearly presented as preliminary hypotheses to be validated by subsequent confirmatory studies. Unfortunately, exploratory results are frequently presented with stronger claims than warranted, leading to overstated conclusions and elevated type I error risks. Therefore, even in exploratory trials, type I error risks should be explicitly acknowledged, and transparent reporting of FWER is advisable [9].

When to adjust for multiple testing

Multiplicity adjustments are critical in studies with: (i) multiple endpoints (e.g. for example, three measures of cardiovascular outcomes); (ii) repeated measures over time (e.g. at three, six, and 12 months); or (iii) multiple treatment arms (e.g. different regimens compared to a shared control arm) [9]. Adjustments are necessary when findings are interpreted collectively for clinical recommendations. For example, in a trial evaluating different schedules or doses of treatment compared to shared control, each experimental arm is part of a family of comparisons and contributes to a broader question, indicating the necessity for multiplicity adjustments [10]. However, adjustments may not be needed if distinct hypotheses inform separate claims of effectiveness. Additional considerations include trials addressing controversial topics or those involving multiple treatments from the same manufacturer, where uncorrected multiplicity may bias conclusions [10].

Multiple endpoints

In trials with coprimary endpoints, where success depends on demonstrating treatment effects across all outcomes, type I error is not inflated, and multiplicity adjustments are unnecessary. These suggestions are consistent with regulatory guidelines for trials aiming for marketing authorization. However, adjustments are required when trials allow

multiple pathways to success—where efficacy in any one outcome is sufficient [11].

Related treatments typically demand adjustments for multiple treatment groups, while unrelated treatments may not [9, 10]. Strategies to mitigate type I errors include prioritizing a single primary outcome, using composite outcomes (though challenging due to varying importance to patients), or applying global tests that sum standardized effect sizes. The win ratio approach offers an alternative by prioritizing critical components, such as cause-specific mortality, and calculating effectiveness through patient-pair comparisons [10, 12].

Adjustments for multiple testing in contemporary clinical trials

Despite established guidelines, implementing multiple-testing adjustments in clinical trials remains suboptimal. A 2012 review of multi-arm trials found that only 62% of studies requiring adjustments accounted for multiplicity, with 38% using ordered comparisons and 18% employing single-step procedures, such as the Bonferroni method. Alarming, 15% of trials with planned adjustments failed to report them in final publications, suggesting selective reporting [13]. Our review of RCTs and high-profile publications (2010–22) confirms the inconsistent use of correction methods across disciplines, which contributes to inflated Type I error rates and irreproducible findings (Table 1).

Multiplicity adjustments remain inconsistently applied across many research areas. In specific fields, including neurology and psychiatry, multiple primary outcomes are necessary to evaluate the effectiveness of an intervention, as long-term mental health conditions require more than one outcome to characterize the effects of treatment sufficiently. Of 55 RCTs on depression (2007–8), only 5.8% accounted for multiplicity, despite most trials reporting multiple primary and secondary outcomes [14]. Similarly, in a review of 209 neurology and psychiatry RCTs (2011–14), 29% involved multiple primary outcomes, yet 75% lacked adjustments; the Bonferroni correction was the most common method when applied [15]. The review was based on clinical trials published in high-impact journals (*The New England Journal of Medicine*, *The Lancet*, *The American Journal of Psychiatry*, *JAMA Psychiatry*, *The Lancet Neurology*, and *Neurology*) and had multiplicity has been addressed, some studies would not remain significant [15].

Trials in pain research frequently neglect adjustments. A 2014 review of 161 RCTs revealed that only 52% specified a primary analysis, and 45% applied corrections. None of the articles that neglected to adjust for multiple analyses acknowledged it in their Methods, Results, or Discussion sections. The 15 studies that reported statistical adjustments used mainly the Holm ($n = 3$ studies), Bonferroni ($n = 2$), and Hochberg ($n = 2$) methods [16]. A systemic review of non-pharmacological pain trials with multiple comparisons found that only 21% of RCTs employed statistical adjustments, primarily using the Bonferroni correction [17].

A systematic review of 388 surgical RCTs (2008–20) identified multiplicity in 175 trials. Adjustments were performed in only 20%, primarily using Bonferroni and Tukey methods. Reporting bias affected 51.7% of studies, undermining reliability [18].

Multiplicity is especially critical in imaging, where studies test dozens of hypotheses. A systematic review of PET/CT

Table 1. Reviews investigating the prevalence of adjustments for multiple comparisons in clinical and epidemiological studies

Study (first author)	Year published	Study interval	Scientific field	Number of studies investigated	Number of studies with multiple comparisons	Proportion of studies with adjustments for multiple testing	Most common method	Journals
Wason <i>et al.</i>	2014	2012	Multiarm clinical trials	59		51%	Hierarchical/closed and Bonferroni	<i>British Medical Journal</i> , <i>The Lancet</i> , <i>New England Journal of Medicine</i> , and <i>PLoS Medicine</i> Six medical journals
Tyler <i>et al.</i>	2011	January 2007 and October 2008	Neurology and psychiatry, clinical trials	55	nearly half (52 reported ≥ 2 primary or secondary outcomes)	3/52 = 5.8%	Bonferroni	
Vickerstaff <i>et al.</i>	2015	July 2011–June 2014	Neurology and psychiatry	209	60 (29% reported multiple primary outcomes)	15/60 = 25%	Bonferroni (used in 6 of 15 adjusted trials)	<i>The New England Journal of Medicine</i> , <i>The Lancet</i> , <i>The American Journal of Psychiatry</i> , <i>JAMA</i> , <i>Neurology</i> , <i>The Lancet Neurology</i> , and <i>The Laryngoscope</i> ; <i>Archives of Otolaryngology—Head & Neck Surgery</i> (now called <i>JAMA Otolaryngology—Head & Neck Surgery</i>); <i>Otolaryngology—Head & Neck Surgery</i> ; and <i>Annals of Otolaryngology, Rhinology, and Laryngology</i> .
Kirkham <i>et al.</i>	2015	2012	Otolaryngology, original human subjects research with samples of >100 subjects	195	140	10%	Bonferroni (used in 5 of 8 corrected studies)	<i>European Journal of Pain</i> , <i>the Journal of Pain</i> , and <i>Pain</i> .
Stacey <i>et al.</i>	2012	2010	Ophthalmology, abstracts presented at ARVO 2010	6415 abstracts	538 (with more than five hypotheses)	14%	Bonferroni and Tukey	<i>European Journal of Pain</i> , <i>the Journal of Pain</i> , and <i>Pain</i> .
Dworkin <i>et al.</i>	2016	January 2006 and June 2013	Pain, randomized clinical trial	101	29	21%	Bonferroni, gatekeeping, Sidák	<i>The New England Journal of Medicine</i> , <i>The Lancet</i> and <i>Circulation</i> . MEDLINE
Gewandter	2014	2006 and 2012	Pain, randomized clinical trial	160	33	45%	Unspecified	
Chalkidou Brand	2014–2021	2000–13 2015–16	PET CT scan Cardiovascular, randomized clinical trial	15 130	15 89 studies with subgroups	1 study 2 studies altogether	Unspecified	
Nevens	2022	2014–19	Pragmatic clinical trials	262 final reports, 159 protocols	38 final reports, 30 protocols	11% final reports, 7% protocols	Bonferroni	

(continued)

Table 1. (continued)

Study (first author)	Year published	Study interval	Scientific field	Number of studies investigated	Number of studies with multiple comparisons	Proportion of studies with adjustments for multiple testing	Most common method	Journals
Pike	2022	January to June 2018	Randomized clinical trial, general medicine	138	28	48% for multiple treatment comparisons (11 out of 28)	Bonferroni, Holm, Hochberg, hierarchical methods	7 journals: <i>Annals of Internal Medicine</i> , <i>British Medical Journal</i> (BMJ), <i>Journal of the American Medical Association</i> (JAMA), <i>Lancet</i> , <i>New England Journal of Medicine</i> (NEJM), <i>National Health Technology Assessment</i> (HTA), or <i>PLOS Medicine</i> (PLOS Med)
Wallach	2018	SATIRE (2007), DISCO (2002–12)	Randomized clinical trial (various medical fields (e. g. cardiovascular, infectious disease))	64 (with at least one subgroup claim in the abstract)	46 (subgroup findings with statistical support)	1/46 = 2.2%	Bonferroni–Holm step-down procedure	
Benjamini	2010	2000–4	Articles, samples from NEJM	60	All had multiplicity issues	22%	Minimal, mostly no methods used	<i>New England Journal of Medicine</i>
Benjamini and Cohen	2017	2000–10	Articles, samples from NEJM	100	All had multiple endpoints	Only 20% addressed multiplicity in any form	Very few used any method	<i>New England Journal of Medicine</i>

biomarker studies (2000–13) found 15 eligible studies, with the number of hypotheses ranging between 8 and 102. Only one study applied adjustments for multiple hypothesis testing, and three studies included validation of their results. The authors estimated an average type-I error probability of 76% (34%–99%), with most published results failing to reach statistical significance [19]. Similarly, unadjusted *P*-values remain an issue in fMRI research: in the Brede database of neuroimaging papers, 49% of studies ($n = 1705$ articles) published uncorrected *P*-values [20].

In ophthalmology, only 14% of abstracts at a major conference with more than five hypothesis tests reported corrections [21]. Similarly, a systematic review of otolaryngology studies found that 72% involved multiple testing, but only 10% addressed it, primarily using Bonferroni corrections. The authors estimated a 41% chance of false positives and noted that 18% of reported *P*-values might be spurious [22].

Multiplicity challenges in pragmatic trials

Pragmatic trials, designed to evaluate real-world effectiveness, often include multiple primary outcomes but rarely apply corrections. Only 10% of such trials (2014–19) reported adjustments. In a review of high-impact journals, inconsistencies were noted: 25% of trials adjusted when all outcomes had to show effectiveness, but only 35% adjusted when not all outcomes needed to succeed. Opinions among statisticians were similarly divided, with adjustments for multiple primary outcomes seen as critical by some but less essential for secondary outcomes or subgroup analyses [23, 24].

Multiplicity challenges in high-impact publications

Even high-profile journals frequently overlook multiplicity adjustments, leading to inflated type I errors. Benjamini and colleagues investigated a batch of papers published in the *New England Journal of Medicine* (NEJM) between 2000 and 2004. Out of the 60 papers, 47 (78.3%) had no multiplicity adjustments, although they should have had in some form [25]. The trend remained, as 80% of papers published between 2000 and 2010 in NEJM investigating multiple endpoints ignored the issue of multiplicity. In multi-arm clinical trials published in 2012 by four top-quality medical journals (*British Medical Journal*, *The Lancet*, *New England Journal of Medicine*, and *PLoS Medicine*), 51% of studies contained adjustment for multiplicity, with a slightly higher proportion of exploratory trials (55%) compared to confirmatory studies (46%). The most common adjustment methods were the gatekeeping/hierarchical/closed approach (24%) and the Bonferroni correction (14%) [9]. A study published in 2020 investigating cardiovascular randomized trials in six high-impact journals found that out of 300 studies with multiplicity, only 28% (85 studies) adjusted the results for multiple comparisons. Interestingly, larger trials were less likely to make adjustments [26].

Multiplicity affecting subgroup analyses

The treatment effect of a new intervention may vary among different segments of the study population. Subgroup analyses may assess the safety profiles and the consistency of the treatment effect across subgroups and detect effects within a subgroup in an otherwise nonsignificant trial. Subgroup analyses may be predefined or a posteriori and may be delineated as confirmatory, exploratory, *post hoc*, and data-driven subgroup identifications [27, 28]. Subgroup analyses carry a

high risk of false positives, especially when numerous comparisons are made. A review of 89 cardiovascular studies showed more common subgroup analyses in trials with non-significant primary results, highlighting potential “fishing for significance.” Only 2% of these trials applied adjustments [29]. A meta-analysis of 64 RCTs with 117 subgroup claims found that only one study used the Bonferroni–Holm procedure, while 39.3% lacked statistical evidence for their claims [30]. Guidelines recommend corrections in subgroup-based comparisons but suggest flexibility when subgroups are nested within treatment arms [31].

Recommended adjustment methods

Effective control of multiplicity is essential to ensure the validity of clinical trial results and prevent erroneous conclusions about treatment efficacy. Adjustment methods must align with the specific clinical and statistical context, considering endpoint priorities, population differences, and correlations between hypotheses. Selecting the optimal approach requires extensive trial simulations to balance sensitivity and specificity.

Multiplicity adjustment strategies in clinical trials generally fall into two extremes: “no adjustment at all” or excessively strict “no error at all” approaches (Table 1). The Bonferroni correction remains the most widely used method for controlling the family-wise error rate (FWER). However, it is often criticized for being overly conservative, increasing the risk of false negatives, and potentially overlooking meaningful effects. Stepwise methods, such as Holm’s procedure (adjusting significance thresholds sequentially) and Hochberg’s method (ranking hypotheses to set critical values), offer a more balanced approach by controlling type I error rates while maintaining statistical power. These methods are increasingly preferred for their ability to balance stringency and flexibility [32].

False discovery rate (FDR) adjustments provide an alternative approach for large-scale studies involving numerous hypotheses, such as genome-wide association studies. FDR-based methods control the proportion of false positives among significant results, yielding *q*-values that balance sensitivity and specificity. These approaches are highly effective for preserving statistical power while minimizing false discoveries [33]. Our prior publication details frequently used adjustment methods, including FWER and FDR-based techniques [32].

Moreover, accessible tools for multiplicity correction offer researchers practical means to maintain statistical rigor. One freely available example is “multipletesting.com,” developed by the authors [32]. This intuitive tool requires no coding skills or specialized expertise, making a multiplicity correction broadly accessible to many users. Additional resources include publicly available FDR adjustment calculators and standard software functions, such as the `p.adjust` function in R.

Gatekeeping procedures handle multiplicity in scenarios where hypotheses are hierarchically or logically related. Primary endpoints are tested first, with secondary endpoints evaluated only if primary tests are significant, safeguarding statistical integrity across multiple comparisons. Methods such as Bonferroni and Holm are common in gatekeeping frameworks, with variants like serial, parallel, and tree-structured approaches tailored for complex trial designs.

Clinical examples show how gatekeeping methods address specific trial design challenges [34, 35].

In drug development, weighted FDR methods account for the relative importance of different endpoints. They achieve greater power by prioritizing primary endpoints and assigning weights to secondary outcomes than traditional hierarchical gatekeeping. These weighted procedures are beneficial in phase II trials where primary endpoints may have low power, enhancing secondary outcome discovery while maintaining rigorous multiplicity control. Such methods apply across diverse research fields where hypotheses vary in importance [36].

Multiplicity in observational epidemiology

Observational epidemiology, which relies on biomarkers or questionnaires to assess exposures, struggles with reproducibility, compounded by high-throughput technologies like transcriptomics and metabolomics [37]. The exposome concept, encompassing all environmental and biological exposures, is vital to understanding gene–environment interactions. Nevertheless, the number of variables in exposome studies often leads to high type I error rates if multiplicity is not addressed [38].

Debate persists over whether multiplicity adjustments are needed in observational studies. Rothman [39] argued against routine adjustments, warning that they could mask important findings by increasing type II errors—particularly in exploratory contexts. However, Rothman's viewpoint assumes researchers transparently present exploratory findings without overstating their significance, which is frequently not the practice case, as transparency regarding multiplicity is often lacking in publications [38]. While genomic studies (e.g. Genome-wide association studies) routinely employ multiplicity adjustments, this is not the norm in broader epidemiology. A 1998 review of articles in the *American Journal of Epidemiology* and the *American Journal of Public Health* revealed a ~20% type I error rate, far exceeding the expected 5% [40].

Nutritional epidemiology particularly exemplifies these challenges. Nonrandomized observational studies in this field often employ flexible statistical adjustments and Food Frequency Questionnaires (FFQs), which can lead to questionable validity. This practice has resulted in almost every food ingredient being linked to cancer risk at some point, despite most findings being implausible or irreproducible [41]. Small effect sizes, confounding, measurement errors, and unadjusted multiplicity exacerbate this issue. Critics argue that focusing on single ingredients in observational or small-scale randomized trials will unlikely advance the field [42]. Instead, large-scale trials evaluating dietary patterns, such as the PREDIMED trial on the Mediterranean diet, offer a more reliable approach to understanding complex nutritional impacts [43].

Transparency in observational studies can be enhanced by explicitly reporting every planned comparison, including measured exposures, outcomes, and their inter-relationships. While the actual number of tests performed can be challenging to capture—especially in exploratory contexts—researchers should strive to track all analyses attempted, even those ultimately excluded from the final report [44]. Preregistration of a Statistical Analysis Plan (SAP) further strengthens transparency by clearly distinguishing confirmatory from exploratory hypotheses, mitigating inflated effect sizes, and selective reporting (as well as HARKing—hypothesizing after

results are known) [45]. Recent metascientific evidence strongly supports the use of preregistration and Registered Reports, confirming that these practices effectively mitigate systematic and publication bias by enabling peers to critically evaluate how rigorously scientific claims have been tested [46]. Despite occasional criticism—such as preregistration discouraging exploratory research or inadvertently signaling study quality without proper scrutiny—these concerns lack strong empirical backing and often arise from misunderstandings about the goal of preregistration [46]. Thus, preregistration addresses Rothman's concerns by transparently delineating exploratory findings without necessarily discouraging exploration itself.

Multiplicity affecting meta-analyses

Systematic meta-analyses effectively summarize accumulated knowledge, aiding clinicians and policymakers. However, multiplicity issues must be addressed explicitly. Detailed prespecification of decision rules for managing multiple groups, time points, and analyses in meta-analysis protocols prevents selective reporting or cherry-picking significant findings from primary studies. Observational studies, often exploratory by nature, are particularly prone to selective inclusion and inflated biases. Unfortunately, observational studies with multiple comparisons, conducted without corrections for multiple hypotheses, will produce meta-analyses with unreliable results [37]. Meta-analyses based on papers with a risk of containing false positives should be treated with care until the credibility of the underlying primary studies is evaluated or further confirmatory studies are conducted.

Conclusions

The reproducibility crisis highlights the urgent need for rigorous statistical practices. Even RCTs, the methodological gold standard, are susceptible to multiplicity issues. Clear prespecification of analytical methods, proper correction for multiplicity, and transparent reporting are essential. Adopting these practices can minimize false-positive findings, enhance reproducibility, and provide a more informed basis for evidence-based clinical and policy decisions.

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Author contributions

Both authors contributed to the conception and design of the study. Data collection and analysis were performed by O.M. The first draft of the manuscript was written by O.M., and both authors commented on previous versions of the manuscript. All authors read and approved the final manuscript. O.M. will act as a guarantor of the paper.

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Data availability

NA.

Use of artificial intelligence (AI) tools

AI was utilized to enhance grammar during the writing process.

Ethics approval

NA.

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